# <span id="page-0-0"></span>Package: estimateW (via r-universe)

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Type Package Title Estimation of Spatial Weight Matrices Version 0.0.1 Author Tamas Krisztin [aut, cre] (<<https://orcid.org/0000-0002-9241-8628>>), Philipp Piribauer [aut] (<<https://orcid.org/0000-0002-3414-3101>>) Maintainer Tamas Krisztin <krisztin@iiasa.ac.at> Description Bayesian estimation of spatial weight matrices in spatial econometric panel models. Allows for estimation of spatial autoregressive (SAR), spatial Durbin (SDM), and spatially lagged explanatory variable (SLX) type specifications featuring an unknown spatial weight matrix. Methodological details are given in Krisztin and Piribauer (2022) [<doi:10.1080/17421772.2022.2095426>](https://doi.org/10.1080/17421772.2022.2095426). License GPL  $(>= 3)$ **Depends** R  $(>= 3.5.0)$ Encoding UTF-8 Language EN-US LazyData true Imports Matrix, matrixcalc, plot.matrix, stats, R6, Rcpp RoxygenNote 7.2.3 LinkingTo Rcpp, RcppArmadillo Repository https://tkrisztin.r-universe.dev RemoteUrl https://github.com/tkrisztin/estimatew RemoteRef HEAD RemoteSha ecc0f625aeb6c73b6fe2a1d76ae06d46b69360ad 1

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<span id="page-1-1"></span>bbinompdf *Probability density for a hierarchical prior setup for the elements of the adjacency matrix based on the beta binomial distribution*

# Description

A hierarchical prior setup can be used in [W\\_priors](#page-35-1) to anchor the prior number of expected neighbors. Assuming a *fixed* prior inclusion probability  $p = 1/2$  for the off-diagonal entries in the binary n by n adjacency matrix  $\Omega$  implies that the number of neighbors (i.e. the row sums of  $\Omega$ ) follows a Binomial distribution with a prior expected number of neighbors for the  $n$  spatial observations of  $(n - 1)p$ . However, such a prior structure has the potential undesirable effect of promoting a relatively large number of neighbors. To put more prior weight on parsimonious neighborhood structures and promote sparsity in  $\Omega$ , the beta binomial prior accounts for the number of neighbors in each row of  $\Omega$ .

# bbinompdf 3

#### Usage

bbinompdf(x, nsize, a, b, min\_k =  $0$ , max\_k = nsize)

# Arguments



#### Details

The beta-binomial distribution is the result of treating the prior inclusion probability  $p$  as random (rather than being fixed) by placing a hierarchical beta prior on it. For the number of neighbors  $x$ , the resulting prior on the elements of  $\Omega$ ,  $\omega_{ij}$ , can be written as:

$$
p(\omega_{ij} = 1|x) \propto \Gamma(a+x) \Gamma(b + (n-1) - x),
$$

where  $\Gamma(\cdot)$  is the Gamma function, and a and b are hyperparameters from the beta prior. In the case of  $a = b = 1$ , the prior takes the form of a discrete uniform distribution over the number of neighbors. By fixing  $a = 1$  the prior can be anchored around the expected number of neighbors m through  $b = [(n-1) - m]/m$  (see Ley and Steel, 2009).

The prior can be truncated by setting a minimum (min\_k) and/or a maximum number of neighbors (max\_k). Values outside this range have zero prior support.

#### Value

Prior density evaluated at x.

#### References

Ley, E., & Steel, M. F. (2009). On the effect of prior assumptions in Bayesian model averaging with applications to growth regression. *Journal of Applied Econometrics*, 24(4). [doi:10.1002/jae.1057.](https://doi.org/10.1002/jae.1057)

<span id="page-3-1"></span><span id="page-3-0"></span>

A four-parameter Beta specification as the prior for the spatial autoregressive parameter  $\rho$ , as proposed by LeSage and Parent (2007) .

#### Usage

betapdf(rho,  $a = 1$ ,  $b = 1$ , rmin = 0, rmax = 1)

# Arguments



#### Details

The prior density is given by:

$$
p(\rho) \sim \frac{1}{Beta(a,b)} \frac{(\rho - \underline{\rho}_{min})^{(a-1)} (\underline{\rho}_{max} - \rho)^{(b-1)}}{2^{a+b-1}}
$$

where  $Beta(a, b)$   $(a, b > 0)$  represents the Beta function,  $Beta(a, b) = \int_0^1 t^{a-1} (1-t)^{b-1} dt$ .

#### Value

Density value evaluated at rho.

# References

LeSage, J. P., and Parent, O. (2007) Bayesian model averaging for spatial econometric models. *Geographical Analysis*, 39(3), 241-267.

<span id="page-4-1"></span><span id="page-4-0"></span>

This function allows the user to specify custom values for Gaussian priors on the slope parameters.

#### Usage

```
beta_priors(
 k,
 beta_mean_prior = matrix(0, k, 1),
 beta_var_prior = diag(k) * 100)
```
# Arguments



# Details

For the slope parameters  $\beta$  the package uses common Normal prior specifications. Specifically,  $p(\beta) \sim \mathcal{N}(\underline{\mu}_{\beta}, \underline{V}_{\beta}).$ 

This function allows the user to specify custom values for the prior hyperparameters  $\mu_{\beta}$  and  $V_{\beta}$ . The default values correspond to weakly informative Gaussian priors with mean zero and a diagonal prior variance-covariance matrix with 100 on the main diagonal.

# Value

A list with the prior mean vector (beta\_mean\_prior), the prior variance matrix (beta\_var\_prior) and the inverse of the prior variance matrix (beta\_var\_prior\_inv).

This class samples slope parameters with a Gaussian prior from the conditional posterior. Use the [beta\\_priors](#page-4-1) class for setup.

#### Format

An [R6Class](#page-0-0) generator object

#### Public fields

beta\_prior The current [beta\\_priors](#page-4-1)

curr\_beta The current value of  $\beta$ 

#### Methods

#### Public methods:

- [beta\\_sampler\\$new\(\)](#page-5-1)
- [beta\\_sampler\\$sample\(\)](#page-5-2)

#### <span id="page-5-1"></span>Method new():

*Usage:*

beta\_sampler\$new(beta\_prior)

*Arguments:*

beta\_prior The list returned by [beta\\_priors](#page-4-1)

# <span id="page-5-2"></span>Method sample():

*Usage:*

beta\_sampler\$sample(Y, X, curr\_sigma)

*Arguments:*

 $Y$  The  $N$  by 1 matrix of responses

X The  $N$  by  $k$  design matrix

curr\_sigma The variance parameter  $\sigma^2$ 

<span id="page-6-0"></span>COVID-19 data set provided by Johns Hopkins University (Dong et al., 2020). The database contains information on (official) daily infections for a large panel of countries around the globe in the very beginning of the outbreak from 17 February to 20 April 2020.

#### Usage

covid

#### Format

A data.frame object.

#### Details

Data is provided for countries: Australia (AUS), Bahrain (BHR), Belgium (BEL), Canada (CAN), China (CHN), Finland (FIN), France (FRA), Germany (DEU), Iran (IRN), Iraq (IRQ), Israel (ISR), Italy (ITA), Japan (JPN), Kuwait (KWT), Lebanon (LBN), Malaysia (MYS), Oman (OMN), Republic of Korea (KOR), Russian Federation (RUS), Singapore (SGP), Spain (ESP), Sweden (SWE), Thailand (THA), United Arab Emirates (ARE), United Kingdom (GBR), United States of America (USA), and Viet Nam (VNM).

The dataset includes daily data on the country specific maximum measured temperature (Temperature) and precipitation levels (Precipitation) as additional covariates (source: Dark Sky API). The stringency index (Stringency) put forward by Hale et al. (2020), which summarizes country-specific governmental policy measures to contain the spread of the virus. We use the biweekly average of the reported stringency index.

#### References

Dong, E., Du, H., and Gardner, L. (2020). An interactive web-based dashboard to track COVID-19 in real time. *The Lancet Infectious Diseases*, 20(5), 533–534. [doi:10.1016/S14733099\(20\)301201.](https://doi.org/10.1016/S1473-3099%2820%2930120-1)

Hale, T., Petherick, A., Phillips, T., and Webster, S. (2020). Variation in government responses to COVID-19. Blavatnik School of Government Working Paper, 31, 2020–2011. [doi:10.1038/s41562-](https://doi.org/10.1038/s41562-021-01079-8) [021010798.](https://doi.org/10.1038/s41562-021-01079-8)

Krisztin, T., and Piribauer, P. (2022). A Bayesian approach for the estimation of weight matrices in spatial autoregressive models, *Spatial Economic Analysis*, 1-20. [doi:10.1080/17421772.2022.2095426.](https://doi.org/10.1080/17421772.2022.2095426)

Krisztin, T., Piribauer, P., and Wögerer, M. (2020). The spatial econometrics of the coronavirus pandemic. *Letters in Spatial and Resource Sciences*, 13 (3), 209-218. [doi:10.1007/s12076020-](https://doi.org/10.1007/s12076-020-00254-1) [002541.](https://doi.org/10.1007/s12076-020-00254-1)

Dong, E., Du, H., and Gardner, L. (2020). An interactive web-based dashboard to track COVID-19 in real time. *The Lancet Infectious Diseases*, 20(5), 533–534. [doi:10.1016/S14733099\(20\)301201.](https://doi.org/10.1016/S1473-3099%2820%2930120-1)

<span id="page-7-0"></span>

While updating the elements of the spatial weight matrix in SAR and SDM type spatial models in a MCMC sampler, the log-determinant term has to be regularly updated, too. When the binary elements of the adjacency matrix are treated unknown, the Matrix Determinant Lemma and the Sherman-Morrison formula are used for computationally efficient updates.

#### Usage

logdetAinvUpdate(ch\_ind, diff, AI, logdet)

#### Arguments



#### Details

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Let  $A = (I_n - \rho W)$  be an invertible n by n matrix. v is an n by 1 column vector of real numbers and  $u$  is a binary vector containing a single one and zeros otherwise. Then the Matrix Determinant Lemma states that:

$$
A + uv' = (1 + v'A^{-1}u)det(A)
$$

This provides an update to the determinant, but the inverse of A has to be updated as well. The Sherman-Morrison formula proves useful:

$$
(A + uv')^{-1} = A^{-1} \frac{A^{-1}uv'A^{-1}}{1 + v'A^{-1}u}
$$

Using these two formulas, an efficient update of the spatial projection matrix determinant can be achieved.

#### Value

A list containing the updated n by n matrix  $A^{-1}$ , as well as the updated log determinant of A

# <span id="page-8-0"></span>logdetPaceBarry 9

#### References

Sherman, J., and Morrison, W. J. (1950) Adjustment of an inverse matrix corresponding to a change in one element of a given matrix. *The Annals of Mathematical Statistics*, 21(1), 124-127.

Harville, D. A. (1998) Matrix algebra from a statistician's perspective. Taylor & Francis.

logdetPaceBarry *Pace and Barry's log determinant approximation*

#### Description

Bayesian estimates of parameters of SAR and SDM type spatial models require the computation of the log-determinant of positive-definite spatial projection matrices of the form  $(I_n - \rho W)$ , where W is a n by n spatial weight matrix. However, direct computation of the log-determinant is computationally expensive.

# Usage

```
logdetPaceBarry(W, length.out = 200, rmin = -1, rmax = 1)
```
#### Arguments



# Details

This function wraps the log-determinant approximation by Barry and Pace (1999), which can be used to precompute the log-determinants over a grid of  $\rho$  values.

#### Value

numeric length.out by 2 matrix; the first column contains the approximated log-determinants the second column the  $\rho$  values ranging between rmin and rmax.

#### References

Barry, R. P., and Pace, R. K. (1999) Monte Carlo estimates of the log determinant of large sparse matrices. *Linear Algebra and its applications*, 289(1-3), 41-54.

<span id="page-9-0"></span>

The sampler uses independent Normal-inverse-Gamma priors to estimate a linear panel data model. The function is used for an illustration on using the [beta\\_sampler](#page-5-3) and [sigma\\_sampler](#page-30-1) classes.

#### Usage

```
normalgamma(
  Y,
  tt,
 X = matrix(1, nrow(Y), 1),niter = 200,
 nretain = 100,
 beta_prior = beta_priors(k = \text{ncol}(X)),
  sigma_prior = sigma_priors()
)
```
# Arguments



# Details

The considered model takes the form:

$$
Y_t = X_t \beta + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, I_n \sigma^2)$ .

<span id="page-10-0"></span> $Y_t$  ( $n \times 1$ ) collects the n cross-sectional observations for time  $t = 1, ..., T$ .  $X_t$  ( $n \times k_1$ ) is a matrix of explanatory variables.  $\beta$  ( $k_1 \times 1$ ) is an unknown slope parameter matrix.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), X = [X'_1, ..., X'_T]'$   $(N \times k)$ , with  $N = nT$ , the final model can be expressed as:

$$
Y = X\beta + \varepsilon,
$$

where  $\varepsilon \sim N(0, I_N \sigma^2)$ . Note that the input data matrices have to be ordered first by the crosssectional (spatial) units and then stacked by time.

#### Examples

```
n = 20; tt = 10; k = 3
X = matrix(stats::rnorm(n*ttk),n*t,t)Y = X % * % c(1,0,-1) + stats::rnorm(n*t,0,.5)res = normalgamma(Y,tt,X)
```
plot.estimateW *Graphical summary of the estimated adjacency matrix* Ω

#### Description

Graphical plot of the posterior probabilities of the estimated adjacency matrix  $\Omega$ .

#### Usage

```
## S3 method for class 'estimateW'
plot(
  x,
  cols = c("white", "lightgrey", "black"),
  breaks = c(0, 0.5, 0.75, 1),
  ...
\mathcal{L}
```


<span id="page-11-0"></span>

Graphical summary of a generated spatial weight matrix

# Usage

## S3 method for class 'sim\_dgp'  $plot(x, \ldots)$ 

# Arguments



<span id="page-11-1"></span>

# Description

Specify prior for the spatial autoregressive parameter and sampling settings

# Usage

```
rho_priors(
 rho_a-prior = 1,
 rho_b_prior = 1,
 rho\_min = 0,
  rho_max = 1,
  init_rho_scale = 1,
 griddy_n = 60,use_griddy_gibbs = TRUE,
 mh_tune_low = 0.4,
 mh_tune_high = 0.6,
 mh\_tune\_scale = 0.1\mathcal{E}
```
# <span id="page-12-0"></span>rho\_sampler 13

# Arguments



<span id="page-12-1"></span>rho\_sampler *An R6 class for sampling the spatial autoregressive parameter* ρ

# Description

An R6 class for sampling the spatial autoregressive parameter  $\rho$ 

An R6 class for sampling the spatial autoregressive parameter  $\rho$ 

# Format

An [R6Class](#page-0-0) generator object

# Details

This class samples the spatial autoregressive parameter using either a tuned random-walk Metropolis-Hastings or a griddy Gibbs step. Use the [rho\\_priors](#page-11-1) class for setup.

For the Griddy-Gibbs algorithm see Ritter and Tanner (1992).

# <span id="page-13-4"></span>Public fields

rho\_prior The current [rho\\_priors](#page-11-1)

curr\_rho The current value of  $\rho$ 

curr\_W The current spatial weight matrix  $W$ ; an n by n matrix.

curr\_A The current spatial filter matrix  $I - \rho W$ .

curr\_AI The inverse of curr\_A

curr\_logdet The current log-determinant of curr\_A

curr\_logdets A set of log-determinants for various values of  $\rho$ . See the [rho\\_priors](#page-11-1) function for settings of step site and other parameters of the grid.

#### Methods

# Public methods:

- [rho\\_sampler\\$new\(\)](#page-13-0)
- [rho\\_sampler\\$stopMHtune\(\)](#page-13-1)
- [rho\\_sampler\\$setW\(\)](#page-13-2)
- [rho\\_sampler\\$sample\(\)](#page-13-3)
- [rho\\_sampler\\$sample\\_Griddy\(\)](#page-14-1)
- [rho\\_sampler\\$sample\\_MH\(\)](#page-14-2)

#### <span id="page-13-0"></span>Method new():

*Usage:* rho\_sampler\$new(rho\_prior, W = NULL)

*Arguments:*

rho\_prior The list returned by [rho\\_priors](#page-11-1)

 $W$  An optional starting value for the spatial weight matrix  $W$ 

<span id="page-13-1"></span>Method stopMHtune(): Function to stop the tuning of the Metropolis-Hastings step. The tuning of the Metropolis-Hastings step is usually carried out until half of the burn-in phase. Call this function to turn it off.

*Usage:* rho\_sampler\$stopMHtune()

#### <span id="page-13-2"></span>Method setW():

*Usage:*

rho\_sampler\$setW(newW, newLogdet = NULL, newA = NULL, newAI = NULL)

*Arguments:*

newW The updated spatial weight matrix W.

newLogdet An optional value for the log determinant corresponding to newW and curr\_rho. newA An optional value for the spatial projection matrix using newW and curr\_rho. newAI An optional value for the matrix inverse of newA.

#### <span id="page-13-3"></span>Method sample():

# <span id="page-14-0"></span>sample W\_fast 15

*Usage:* rho\_sampler\$sample(Y, mu, sigma) *Arguments:* Y The  $n$  by  $T$  matrix of responses. mu The  $n$  by  $T$  matrix of means. sigma The variance parameter  $\sigma^2$ .

# <span id="page-14-1"></span>Method sample\_Griddy():

```
Usage:
rho_sampler$sample_Griddy(Y, mu, sigma)
Arguments:
Y The n by T matrix of responses.
mu The n by T matrix of means.
sigma The variance parameter \sigma^2.
```
# <span id="page-14-2"></span>Method sample\_MH():

*Usage:* rho\_sampler\$sample\_MH(Y, mu, sigma)

*Arguments:*

Y The  $n$  by  $T$  matrix of responses. mu The  $n$  by  $T$  matrix of means. sigma The variance parameter  $\sigma^2$ .

# References

Ritter, C., and Tanner, M. A. (1992). Facilitating the Gibbs sampler: The Gibbs stopper and the griddy-Gibbs sampler. *Journal of the American Statistical Association*, 87(419), 861-868.

sampleW\_fast *A fast sampling step implemented in C++ for updating the spatial weight matrix.*

# Description

This function is intended to be called from the R6 class  $W$ \_sampler.

# Usage

```
sampleW_fast(
  Y,
  curr_sigma,
 mu,
  lag_mu,
 W_prior,
```

```
curr_W,
 curr_w,
 curr_A,
 curr_AI,
 curr_logdet,
 curr_rho,
 nr_neighbors_prior,
 symmetric,
 spatial_error,
 row_standardized
\mathcal{L}
```


<span id="page-15-0"></span>

<span id="page-16-0"></span>sar *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial autoregressive model (SAR) with exogenous spatial weight matrix.*

# Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter beta prior for the spatial autoregressive parameter  $\rho$ . The function is used as an illustration on using the [beta\\_sampler](#page-5-3), [sigma\\_sampler](#page-30-1), and [rho\\_sampler](#page-12-1) classes.

# Usage

```
sar(
 Y,
 tt,
 W,
 Z = matrix(1, nrow(Y), 1),niter = 200,nretain = 100,
  rho_prior = rho_priors(),
 beta_prior = beta_priors(k = ncol(Z)),
  sigma_prior = sigma_priors()
\mathcal{L}
```


#### <span id="page-17-0"></span>Details

The considered panel spatial autoregressive model (SAR) takes the form:

$$
Y_t = \rho W Y_t + Z_t \beta + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, I_n \sigma^2)$ . The row-stochastic n by n spatial weight matrix W is non-negative and has zeros on the main diagonal.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $Z_t$  ( $n \times k_3$ ) is a matrix of explanatory variables.  $\beta$  ( $k_3 \times 1$ ) is an unknown slope parameter matrix.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), Z = [Z'_1, ..., Z'_T]'$   $(N \times k_3)$ , with  $N = nT$ , the final model can be expressed as:

$$
Y = \rho \tilde{W}Y + Z\beta + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, I_N \sigma^2)$ . Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time. This is a wrapper function calling [sdm](#page-23-1) with no spatially lagged dependent variables.

#### Examples

```
n = 20; tt = 10
dgp_dat = sim\_dgp(n = n, t = tt, rho = .5, beta3 = c(1,.5), sigma2 = .5)res = sar(Y = dgp_dataY, tt = tt, W = dgp_dataW,Z = dgp_data$Z,niter = 100,nretain = 50
```


sarw *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial autoregressive model (SAR) with unknown spatial weight matrix*

#### Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter beta prior for the spatial autoregressive parameter  $\rho$ . This is a wrapper function calling [sdmw](#page-24-1) with no spatially lagged exogenous variables.

#### Usage

```
sarw(
  Y,
  tt,
  Z,
 niter = 100,
  nretain = 50,
 W_prior = W_priors(n = nrow(Y)/tt),
  rho\_prior = rho\_priors(),
 beta_prior = beta_priors(k = \text{ncol}(Z)),
  sigma_prior = sigma_priors()
)
```
<span id="page-18-0"></span>sarw 19

# Arguments



#### Details

The considered panel spatial autoregressive model (SAR) with unknown  $(n \text{ by } n)$  spatial weight matrix  $W$  takes the form:

$$
Y_t = \rho W Y_t + Z\beta + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, I_n \sigma^2)$  and  $W = f(\Omega)$ . The *n* by *n* matrix  $\Omega$  is an unknown binary adjacency matrix with zeros on the main diagonal and  $f(\cdot)$  is the (optional) row-standardization function.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $Z_t$  ( $n \times k_3$ ) is a matrix of explanatory variables.  $\beta$  ( $k_3 \times 1$ ) is an unknown slope parameter vector.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1)$ , and  $Z = [Z'_1, ..., Z'_T]'$  $(N \times k_3)$ , with  $N = nT$ . The final model can be expressed as:

$$
Y = \rho \tilde{W}Y + Z\beta + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, I_N \sigma^2)$ . Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

Estimation usually even works well in cases of  $n \gg T$ . However, note that for applications with  $n > 200$  the estimation process becomes computationally demanding and slow. Consider in this case reducing niter and nretain and carefully check whether the posterior chains have converged.

# Value

List with posterior samples for the slope parameters,  $\rho$ ,  $\sigma^2$ , W, and average direct, indirect, and total effects.

# Examples

```
n = 20; tt = 10
dgp_dat = sim\_dgp(n = n, tt = tt, rho = .5, beta3 = c(.5,1),sigma2 = .05, n\_neighbor = 3, intercept = TRUEres = sarw(Y = dgp_dat$Y, tt = tt, Z = dgp_dat$Z, niter = 20, nretain = 10)
```


sdem *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial Durbin error model (SDEM) with exogenous spatial weight matrix.*

#### Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter prior for the spatial autoregressive parameter  $\rho$ . The function is used as an illustration on using the [beta\\_sampler](#page-5-3), [sigma\\_sampler](#page-30-1), and [rho\\_sampler](#page-12-1) classes.

#### Usage

```
sdem(
  Y,
  tt,
 W,
  X = matrix(0, nrow(Y), 0),Z = matrix(1, nrow(Y), 1),niter = 200,
  nretain = 100,
  rho_prior = rho_priors(),
 beta_prior = beta_priors(k = \text{ncol}(X) \times 2 + \text{ncol}(Z)),
  sigma_prior = sigma_priors()
)
```


<span id="page-19-0"></span>

<span id="page-20-0"></span>sdem 21



# Details

The considered panel spatial Durbin error model (SDEM) takes the form:

$$
Y_t = X_t \beta_1 + W X_t \beta_2 + Z \beta_3 + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, (I_n - \rho W)\sigma^2)$ . The row-stochastic n by n spatial weight matrix W is non-negative and has zeros on the main diagonal.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $X_t$  ( $n \times k_1$ ) and  $Z_t$  ( $n \times k_2$ ) are matrices of explanatory variables, where the former will also be spatially lagged.  $\beta_1$  ( $k_1 \times 1$ ),  $\beta_2$  ( $k_1 \times 1$ ) and  $\beta_3$  ( $k_2 \times 1$ ) are unknown slope parameter vectors.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), X = [X'_1, ..., X'_T]'$   $(N \times k_1)$ and  $Z = [Z'_1, ..., Z'_T]'$   $(N \times k_2)$ , with  $N = nT$ , the final model can be expressed as:

$$
Y = X\beta_1 + \tilde{W}X\beta_2 + Z\beta_3 + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, \sigma^2(I_n \otimes (I_n - \rho W)).$  Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

#### Examples

```
n = 20; tt = 10
dgp_dat = sim_dgp(n = n, tt = tt, rho = .5, beta1 = c(.5,1), beta2 = c(-1,.5),
                   beta3 = c(1.5), sigma2 = .5, spatial_error = TRUE)
res = sdem(Y = dgp_dataY, tt = tt, W = dgp_dataW, X = dgp_dataX,Z = dgp_data$Z, niter = 100, nretain = 50)
```
<span id="page-21-1"></span><span id="page-21-0"></span>sdemw *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial Durbin error model (SDEM) with unknown spatial weight matrix*

# Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter beta prior for the spatial autoregressive parameter  $\rho$ . It is a wrapper around [W\\_sampler](#page-35-2).

# Usage

```
sdemw(
  Y,
  tt,
 X = matrix(\emptyset, nrow(Y), \emptyset),Z = matrix(1, nrow(Y), 1),niter = 100,
 nretain = 50,
 W_prior = W_priors(n = nrow(Y)/tt),
  rho_prior = rho_priors(),
 beta_prior = beta_priors(k = \text{ncol}(X) * 2 + \text{ncol}(Z)),
  sigma_prior = sigma_priors()
)
```


#### <span id="page-22-0"></span>sdemw 23



#### Details

The considered panel spatial Durbin error model (SDEM) with unknown  $(n \text{ by } n)$  spatial weight matrix  $W$  takes the form:

$$
Y_t = X_t \beta_1 + W X_t \beta_2 + Z \beta_3 + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, (I_n - \rho W)\sigma^2)$  and  $W = f(\Omega)$ . The *n* by *n* matrix  $\Omega$  is an unknown binary adjacency matrix with zeros on the main diagonal and  $f(\cdot)$  is the (optional) row-standardization function.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $X_t$  ( $n \times k_1$ ) and  $Z_t$  ( $n \times k_2$ ) are matrices of explanatory variables, where the former will also be spatially lagged.  $\beta_1$  ( $k_1 \times 1$ ),  $\beta_2$  ( $k_1 \times 1$ ) and  $\beta_3$  ( $k_2 \times 1$ ) are unknown slope parameter vectors.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), X = [X'_1, ..., X'_T]'$   $(N \times k_1)$ and  $Z = [Z'_1, ..., Z'_T]'$   $(N \times k_2)$ , with  $N = nT$ . The final model can be expressed as:

$$
Y = X\beta_1 + \tilde{W}X\beta_2 + Z\beta_3 + \varepsilon,
$$

where  $\varepsilon \sim N(0, \sigma^2(I_n \otimes (I_n - \rho W)).$  Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

Estimation usually even works well in cases of  $n >> T$ . However, note that for applications with  $n > 200$  the estimation process becomes computationally demanding and slow. Consider in this case reducing niter and nretain and carefully check whether the posterior chains have converged.

#### Value

List with posterior samples for the slope parameters,  $\rho$ ,  $\sigma^2$ , W, and average direct, indirect, and total effects.

#### Examples

```
n = 20; tt = 10
dgp_dat = sim_dgp(n =n, tt = tt, rho = .75, beta1 = c(.5,1), beta2 = c(-1,.5),
        beta3 = c(1.5), sigma2 = .05,n_neighbor = 3,intercept = TRUE, spatial_error = TRUE)
# res = sdemw(Y = dgp_dat$Y,tt = tt,X = dgp_dat$X,Z = dgp_dat$Z,niter = 20,nretain = 10)
```
<span id="page-23-1"></span><span id="page-23-0"></span>sdm *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial Durbin model (SDM) with exogenous spatial weight matrix.*

# Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter prior for the spatial autoregressive parameter  $\rho$ . The function is used as an illustration on using the [beta\\_sampler](#page-5-3), [sigma\\_sampler](#page-30-1), and [rho\\_sampler](#page-12-1) classes.

# Usage

```
sdm(
  Y,
 tt,
 W,
 X = matrix(\emptyset, nrow(Y), \emptyset),Z = matrix(1, nrow(Y), 1),niter = 200,
  nretain = 100,
  rho_prior = rho_priors(),
  beta_prior = beta_priors(k = \text{ncol}(X) * 2 + \text{ncol}(Z)),
  sigma_prior = sigma_priors()
)
```


#### <span id="page-24-0"></span>sdmw 25



#### Details

The considered panel spatial Durbin model (SDM) takes the form:

$$
Y_t = \rho W Y_t + X_t \beta_1 + W X_t \beta_2 + Z \beta_3 + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, I_n \sigma^2)$ . The row-stochastic n by n spatial weight matrix W is non-negative and has zeros on the main diagonal.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $X_t$  ( $n \times k_1$ ) and  $Z_t$  ( $n \times k_2$ ) are matrices of explanatory variables, where the former will also be spatially lagged.  $\beta_1$  ( $k_1 \times 1$ ),  $\beta_2$  ( $k_1 \times 1$ ) and  $\beta_3$  ( $k_2 \times 1$ ) are unknown slope parameter vectors.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), X = [X'_1, ..., X'_T]'$   $(N \times k_1)$ and  $Z = [Z'_1, ..., Z'_T]'$   $(N \times k_2)$ , with  $N = nT$ , the final model can be expressed as:

$$
Y = \rho \tilde{W}Y + X\beta_1 + \tilde{W}X\beta_2 + Z\beta_3 + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, I_N \sigma^2)$ . Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

#### Examples

```
n = 20; tt = 10
dgp_dat = \sin\phigp(n = n, tt = tt, rho = .5, beta1 = c(.5,1), beta2 = c(-1,.5),
                  beta3 = c(1.5), sigma2 = .5)
res = sdm(Y = dgp_dataSY, tt = tt, W = dgp_dataSW, X = dgp_dataSX,Z = dgp_dat$Z, niter = 100, nretain = 50)
```
<span id="page-24-1"></span>

sdmw *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial Durbin model (SDM) with unknown spatial weight matrix*

#### **Description**

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter beta prior for the spatial autoregressive parameter  $\rho$ . It is a wrapper around [W\\_sampler](#page-35-2).

# Usage

```
sdmw(
  Y,
  tt,
  X = matrix(0, nrow(Y), 0),Z = matrix(1, nrow(Y), 1),niter = 100,
  nretain = 50,
  W_prior = W_priors(n = nrow(Y)/tt),
  rho_prior = rho_priors(),
  beta_prior = beta_priors(k = \text{ncol}(X) * 2 + \text{ncol}(Z)),
  sigma_prior = sigma_priors()
\mathcal{L}
```
# Arguments



# Details

The considered panel spatial Durbin model (SDM) with unknown  $(n \text{ by } n)$  spatial weight matrix W takes the form:

<span id="page-25-0"></span>

$$
Y_t = \rho W Y_t + X_t \beta_1 + W X_t \beta_2 + Z \beta_3 + \varepsilon_t,
$$

<span id="page-26-0"></span>with  $\varepsilon_t \sim N(0, I_n \sigma^2)$  and  $W = f(\Omega)$ . The *n* by *n* matrix  $\Omega$  is an unknown binary adjacency matrix with zeros on the main diagonal and  $f(\cdot)$  is the (optional) row-standardization function.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the *n* cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $X_t$  ( $n \times k_1$ ) and  $Z_t$  ( $n \times k_2$ ) are matrices of explanatory variables, where the former will also be spatially lagged.  $\beta_1$  ( $k_1 \times 1$ ),  $\beta_2$  ( $k_1 \times 1$ ) and  $\beta_3$  ( $k_2 \times 1$ ) are unknown slope parameter vectors.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), X = [X'_1, ..., X'_T]'$   $(N \times k_1)$ and  $Z = [Z'_1, ..., Z'_T]'$   $(N \times k_2)$ , with  $N = nT$ . The final model can be expressed as:

$$
Y = \rho \tilde{W}Y + X\beta_1 + \tilde{W}X\beta_2 + Z\beta_3 + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, I_N \sigma^2)$ . Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

Estimation usually even works well in cases of  $n \gg T$ . However, note that for applications with  $n > 200$  the estimation process becomes computationally demanding and slow. Consider in this case reducing niter and nretain and carefully check whether the posterior chains have converged.

#### Value

List with posterior samples for the slope parameters,  $\rho$ ,  $\sigma^2$ , W, and average direct, indirect, and total effects.

#### Examples

```
n = 20; tt = 10
dgp_dat = sim_dgp(n =n, tt = tt, rho = .75, beta1 = c(.5,1), beta2 = c(-1,.5),
            beta3 = c(1.5), sigma2 = .05, n_neighbor = 3, intercept = TRUE)
res = sdmw(Y = dgp_dat$Y,tt = tt,X = dgp_dat$X,Z = dgp_dat$Z,niter = 20,nretain = 10)
```


#### Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter prior for the spatial autoregressive parameter  $\rho$ . The function is used as an illustration on using the [beta\\_sampler](#page-5-3), [sigma\\_sampler](#page-30-1), and [rho\\_sampler](#page-12-1) classes.

# Usage

```
sem(
 Y,
  tt,
 W,
  Z = matrix(1, nrow(Y), 1),niter = 200,
 nretain = 100,
  rho_prior = rho_priors(),
 beta_prior = beta_priors(k = \text{ncol}(Z)),
  sigma_prior = sigma_priors()
)
```
# Arguments



# Details

The considered panel spatial error model (SDEM) takes the form:

$$
Y_t = Z\beta + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, (I_n - \rho W)\sigma^2)$ . The row-stochastic n by n spatial weight matrix W is non-negative and has zeros on the main diagonal.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the *n* cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $Z_t$  ( $n \times k_2$ ) is a matrix of explanatory variables, where the former will also be spatially lagged.  $\beta$  ( $k_3 \times 1$ ) is an unknown slope parameter vector.

<span id="page-27-0"></span>

<span id="page-28-0"></span>semw 29

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1)$  and  $Z = [Z'_1, ..., Z'_T]'$  $(N \times k_2)$ , with  $N = nT$ , the final model can be expressed as:

$$
Y = Z\beta_3 + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, \sigma^2(I_n \otimes (I_n - \rho W)).$  Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

#### Examples

```
n = 20; tt = 10
dgp_dat = \sin\phigp(n =n, tt = tt, rho = .5, beta3 = c(.5,1),
            sigma2 = .05,n_neighbor = 3,intercept = TRUE,spatial_error = TRUE)
res = sem(Y = dgp_dataSY, tt = tt, W = dgp_dataSW,Z = dgp_data$Z, niter = 100, nretain = 50)
```


#### Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters, as well as a four-parameter beta prior for the spatial autoregressive parameter  $\rho$ . This is a wrapper function calling [sdemw](#page-21-1) with no spatially lagged exogenous variables.

#### Usage

```
semw(
 Y,
  tt,
  Z,
 niter = 100,
 nretain = 50,
 W_prior = W_priors(n = nrow(Y)/tt),
  rho_prior = rho_priors(),
 beta\_prior = beta\_priors(k = ncol(Z)),sigma_prior = sigma_priors()
\mathcal{L}
```


<span id="page-29-0"></span>

#### Details

The considered panel spatial error model (SEM) with unknown (n by n) spatial weight matrix W takes the form:

$$
Y_t = Z\beta + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, (I_n - \rho W)\sigma^2)$  and  $W = f(\Omega)$ . The *n* by *n* matrix  $\Omega$  is an unknown binary adjacency matrix with zeros on the main diagonal and  $f(\cdot)$  is the (optional) row-standardization function.  $\rho$  is a scalar spatial autoregressive parameter.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $Z_t$  ( $n \times k_3$ ) is a matrix of explanatory variables.  $\beta$  ( $k_3 \times 1$ ) is an unknown slope parameter vector.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1)$ , and  $Z = [Z'_1, ..., Z'_T]'$  $(N \times k_3)$ , with  $N = nT$ . The final model can be expressed as:

$$
Y = Z\beta + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, \sigma^2(I_n \otimes (I_n - \rho W)).$  Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

Estimation usually even works well in cases of  $n \gg T$ . However, note that for applications with  $n > 200$  the estimation process becomes computationally demanding and slow. Consider in this case reducing niter and nretain and carefully check whether the posterior chains have converged.

#### Value

List with posterior samples for the slope parameters,  $\rho$ ,  $\sigma^2$ , W, and average direct, indirect, and total effects.

#### Examples

```
n = 20; tt = 10
dgp_dat = sim\_dgp(n = n, tt = tt, rho = .5, beta3 = c(.5,1),sigma2 = .05,n_neighbor = 3,intercept = TRUE,spatial_error = TRUE)
res = semw(Y = dgp_dat$Y, tt = tt, Z = dgp_dat$Z, niter = 20, nretain = 10)
```
<span id="page-30-2"></span><span id="page-30-0"></span>

Set prior specification for the error variance using an inverse Gamma distribution

#### Usage

```
sigma_priors(sigma_rate_prior = 0.001, sigma_shape_prior = 0.001)
```
#### Arguments

sigma\_rate\_prior Sigma rate prior parameter (scalar), default: 0.001.

sigma\_shape\_prior

Sigma shape prior parameter (scalar), default: 0.001. This function allows the user to specify priors for the error variance  $\sigma^2$ .

<span id="page-30-1"></span>sigma\_sampler *An R6 class for sampling for sampling* σˆ2

# Description

This class samples nuisance parameter which an inverse Gamma prior from the conditional posterior. Use the [sigma\\_priors](#page-30-2) class for setup.

# Format

An [R6Class](#page-0-0) generator object

#### Public fields

sigma\_prior The current [sigma\\_priors](#page-30-2)

curr\_sigma The current value of  $\sigma^2$ 

#### Methods

Public methods:

- [sigma\\_sampler\\$new\(\)](#page-30-3)
- [sigma\\_sampler\\$sample\(\)](#page-31-1)

<span id="page-30-3"></span>Method new():

*Usage:*

```
sigma_sampler$new(sigma_prior)
 Arguments:
 sigma_prior The list returned by sigma_priors
Method sample():
 Usage:
 sigma_sampler$sample(Y, mu)
 Arguments:
 Y The N by 1 matrix of responses
```
mu The  $N$  by 1 matrix of means

sim\_dgp *Simulating from a data generating process*

# Description

This function can be used to generate data from a data generating process for SDM, SAR, SEM, and SLX type models.

#### Usage

```
sim_dgp(
 n,
  tt,
  rho,
 beta1 = c(),
 beta2 = c(),
 beta3 = c(),
  sigma2,
 n_{\text{neighbour}} = 4,
 W = NULL,do_symmetric = FALSE,
  intercept = FALSE,
  spatial_error = FALSE
)
```




#### Details

For the SDM, SAR, and SLX models the generated spatial panel model takes the form

$$
Y = \rho WY + X\beta_1 + WX\beta_2 + Z\beta_3 + \epsilon,
$$

with  $\epsilon \sim N(0, I_n \sigma^2)$ .

For the SEM model the generated spatial panel model takes the form

$$
Y = X\beta_1 + WX\beta_2 + Z\beta_3 + \epsilon,
$$

with  $\epsilon \sim N(0, (I_n - \rho W)\sigma^2)$ .

The function generates the  $N \times 1$  vector Y. The elements of the explanatory variable matrices X  $(N \times k_1)$  and  $Z (N \times k_2)$  are randomly generated from a Gaussian distribution with zero mean and unity variance  $(N(0, 1))$ .

The non-negative, row-stochastic n by n matrix  $W$  is constructed using a k-nearest neighbor specification based on a randomly generated spatial location pattern, with coordinates sampled from a standard normal distribution.

Values for the parameters  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$ , as well as  $\rho$  and  $\sigma^2$  have to be provided by the user. The length of  $\beta_1$  and  $\beta_2$  have to be equal.

- A spatial Durbin model (SDM) is constructed if  $\rho$  is not equal to zero, spatial error is FALSE, and  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  are all supplied by the user.
- A spatial autoregressive model is constructed if  $\rho$  is not equal to zero, spatial error is FALSE, and only  $\beta_3$  is supplied by the user.
- An SLX type model is constructed if  $\rho$  is equal to zero, spatial error is FALSE, and  $\beta_1$ ,  $\beta_2$ are supplied by the user.
- An SEM type model is constructed if spatial\_error is TRUE and either only  $\beta_3$  or  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  are supplied by the user.

#### Value

A list with the generated  $X, Y$  and  $W$  and a list of parameters.

# Examples

```
# SDM data generating process
dgp_dat = \sin\phigp(n =20, tt = 10, rho = .5, beta1 = c(1, -1),
                  beta2 = c(0,.5), beta3 = c(.2), sigma2 = .5)
```


w *A Markov Chain Monte Carlo (MCMC) sampler for the panel spatial SLX model with unknown spatial weight matrix*

#### Description

The sampler uses independent Normal-inverse-Gamma priors for the slope and variance parameters. It is a wrapper around [W\\_sampler](#page-35-2).

# Usage

```
slxw(
 Y,
 tt,
 X = matrix(0, nrow(Y), 0),Z = matrix(1, nrow(Y), 1),niter = 100,
 nretain = 50,
 W_pprior = W_ppriors(n = nrow(Y)/tt),
 beta_prior = beta_priors(k = ncol(X) * 2 + ncol(Z)),
  sigma_prior = sigma_priors()
\mathcal{L}
```


<span id="page-33-0"></span>

<span id="page-34-0"></span>

#### Details

The considered spatial panel SLX model with unknown  $(n \text{ by } n)$  spatial weight matrix W takes the form:

$$
Y_t = X_t \beta_1 + W X_t \beta_2 + Z \beta_3 + \varepsilon_t,
$$

with  $\varepsilon_t \sim N(0, I_n \sigma^2)$  and  $W = f(\Omega)$ . The *n* by *n* matrix  $\Omega$  is an unknown binary adjacency matrix with zeros on the main diagonal and  $f(\cdot)$  is the (optional) row-standardization function.

 $Y_t$  ( $n \times 1$ ) collects the n cross-sectional (spatial) observations for time  $t = 1, ..., T$ .  $X_t$  ( $n \times k_1$ ) and  $Z_t$  ( $n \times k_2$ ) are matrices of explanatory variables, where the former will also be spatially lagged.  $\beta_1$  ( $k_1 \times 1$ ),  $\beta_2$  ( $k_1 \times 1$ ) and  $\beta_3$  ( $k_2 \times 1$ ) are unknown slope parameter vectors.

After vertically staking the T cross-sections  $Y = [Y'_1, ..., Y'_T]'$   $(N \times 1), X = [X'_1, ..., X'_T]'$   $(N \times k_1)$ and  $Z = [Z'_1, ..., Z'_T]'$   $(N \times k_2)$ , with  $N = nT$ . The final model can be expressed as:

$$
Y = X\beta_1 + \tilde{W}X\beta_2 + Z\beta_3 + \varepsilon,
$$

where  $\tilde{W} = I_T \otimes W$  and  $\varepsilon \sim N(0, I_N \sigma^2)$ . Note that the input data matrices have to be ordered first by the cross-sectional spatial units and then stacked by time.

Estimation usually even works well in cases of  $n \gg T$ . However, note that for applications with  $n > 200$  the estimation process becomes computationally demanding and slow. Consider in this case reducing niter and nretain and carefully check whether the posterior chains have converged.

# Value

List with posterior samples for the slope parameters,  $\sigma^2$ , W, and average direct, indirect, and total effects.

#### Examples

```
set.seed(123)
n = 20; tt = 10
dgp_dat = \sin\phigp(n = 20, tt = 10, rho = 0, beta1 = c(1,-1),
                  beta2 = c(3, -2.5), beta3 = c(.2), sigma2 = .05,
                  n_neighbor = 3,intercept = TRUE)
res = slxw(Y = dgp_data*Y, tt = tt, X = dgp_data*X, Z = dgp_data*Z,niter = 20, nretain = 10)
```
<span id="page-35-1"></span><span id="page-35-0"></span>

Set prior specifications for the n by n spatial weight matrix  $W = f(\Omega)$ , where  $\Omega$  is an n by n unknown binary adjacency matrix (with zeros on the main diagonal), and  $f()$  denotes the (optional) row-standardization function

#### Usage

```
W_priors(
  n,
  W_prior = matrix(0.5, n, n),
  symmetric_prior = FALSE,
  row_standardized_prior = TRUE,
 nr_neighbors_prior = bbinompdf(0:(n - 1), nsize = n - 1, a = 1, b = 1, min_k = 0, max_k
   = n - 1)
```
# Arguments



<span id="page-35-2"></span>W\_sampler *An R6 class for sampling the elements of* W

# Description

An R6 class for sampling the elements of  $W$ An R6 class for sampling the elements of  $W$ 

# <span id="page-36-2"></span>W\_sampler 37

# Format

An [R6Class](#page-0-0) generator object

# Details

This class samples the spatial weight matrix. Use the function [W\\_priors](#page-35-1) class for setup.

The sampling procedure relies on conditional Bernoulli posteriors outlined in Krisztin and Piribauer (2022).

# Public fields

W\_prior The current [W\\_priors](#page-35-1)

- curr\_w numeric, non-negative  $n$  by  $n$  spatial weight matrix with zeros on the main diagonal. Depending on the [W\\_priors](#page-35-1) settings can be symmetric and/or row-standardized.
- curr\_W binary  $n$  by  $n$  spatial connectivity matrix  $\Omega$
- curr\_A The current spatial projection matrix  $I \rho W$ .

curr\_AI The inverse of curr\_A

- curr\_logdet The current log-determinant of curr\_A
- curr\_rho single number between -1 and 1 or NULL, depending on whether the sampler updates the spatial autoregressive parameter  $\rho$ . Set while invoking initialize or using the function set\_rho.

spatial\_error Should a spatial error model be constructed? Defaults to FALSE.

# Methods

#### Public methods:

- [W\\_sampler\\$new\(\)](#page-36-0)
- [W\\_sampler\\$set\\_rho\(\)](#page-36-1)
- [W\\_sampler\\$sample\\_fast\(\)](#page-37-0)
- [W\\_sampler\\$sample\(\)](#page-37-1)

#### <span id="page-36-0"></span>Method new():

*Usage:*

W\_sampler\$new(W\_prior, curr\_rho = NULL, spatial\_error = FALSE)

*Arguments:*

W\_prior The list returned by [W\\_priors](#page-35-1)

- curr\_rho Optional single number between -1 and 1. Value of the spatial autoregressive parameter  $\rho$ . Defaults to NULL, in which case no updates of the log-determinant, the spatial projection matrix, and its inverse are carried out.
- <span id="page-36-1"></span>spatial\_error Optional binary, specifying whether the sampler is for a spatial error model (TRUE) or for a spatial autoregressive specification (FALSE). Defaults to FALSE. If spatial\_error = TRUE then a value curr\_rho has to be supplied at initialization.

Method set\_rho(): If the spatial autoregressive parameter  $\rho$  is updated during the sampling procedure the log determinant, the spatial projection matrix  $I - \rho W$  and it's inverse must be updated. This function should be used for a consistent update. At least the new scalar value for  $\rho$ must be supplied.

*Usage:*

```
W_sampler$set_rho(new_rho, newLogdet = NULL, newA = NULL, newAI = NULL)
```
*Arguments:*

new\_rho single, number; must be between -1 and 1.

newLogdet An optional value for the log determinant corresponding to newW and curr\_rho newA An optional value for the spatial projection matrix using newW and curr\_rho newAI An optional value for the matrix inverse of newA

<span id="page-37-0"></span>Method sample\_fast():

```
Usage:
W_sampler$sample_fast(
 Y,
 curr_sigma,
 mu,
  lag_mu = matrix(0, nrow(tY), ncol(tY)))
```
*Arguments:*

Y The  $n$  by  $tt$  matrix of responses

curr\_sigma The variance parameter  $\sigma^2$ 

mu The  $n$  by  $tt$  matrix of means.

lag\_mu  $n$  by tt matrix of means that will be spatially lagged with the estimated W. Defaults to a matrix with zero elements.

#### <span id="page-37-1"></span>Method sample():

*Usage:*

```
W_sampler$sample(Y, curr_sigma, mu, lag_mu = matrix(0, nrow(tY), ncol(tY)))
```
*Arguments:*

Y The  $n$  by  $tt$  matrix of responses

curr\_sigma The variance parameter  $\sigma^2$ 

mu The  $n$  by  $tt$  matrix of means.

lag\_mu n by tt matrix of means that will be spatially lagged with the estimated W. Defaults to a matrix with zero elements.

# References

Krisztin, T., and Piribauer, P. (2022) A Bayesian approach for the estimation of weight matrices in spatial autoregressive models. *Spatial Economic Analysis*, 1-20.

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